## Estimating CAPM using GMM

## Testing the CAPM in a GMM setting

Consider the CAPM relationship, specified by the "moment condition"

$$
E\left[e r_{i t}\right]=\beta_{i} E\left[e r_{m t}\right]
$$

where $e r_{i t}$ is the excess return on an asset or a portfolio, and $e r_{m t}$ the excess return on the market portfolio.
One can alternatively specify the CAPM more generally as

$$
E\left[r_{i t}\right]=E\left[r_{z t}\right]+\beta_{i}\left(E\left[r_{m t}\right]-E\left[r_{z t}\right]\right)
$$

where we let $r_{z t}$ be the return on a "zero covariance" portfolio.

The typical way of testing this relationship, the Black et al. [1972] method, estimates the corresponding regression

$$
e r_{i t}=\alpha_{i}+\beta_{i} e r_{m t}+\varepsilon_{i t}
$$

and tests whether $\alpha_{i}=0$ on an equation by equation basis.
Econometically wasteful: the restriction on the constant will hold for all assets, one want to test this jointly, in a multivariate setting. Doing this was proposed by Gibbons [1982], who showed how one could construct a multivariate statistic for testing this. His method was later expanded upon by Gibbons et al. [1989].
These statistic were developed under distributional assumptions that allowed us to use Maximum Likelihood, namely multivariate normality.

What if these are not fulfilled, can we still construct a similar test statistic?
This is done in MacKinlay and Richardson [1991] (MR). They construct a test statistic that essentially tests the same restriction, that $\alpha_{I}=0$ or that $\alpha_{i}=E\left[r_{z t}\right]\left(1-\beta_{i}\right)$, but in a GMM framework, not a ML.

The setup:
The usual regression

$$
r_{i t}=\alpha_{i}+\beta_{i} r_{m t}+\varepsilon_{i t}
$$

Assume that

$$
E\left[\varepsilon_{i t} \mid r_{m t}\right]=0
$$

Two moment restrictions for each asset $i$ :

$$
\begin{aligned}
& E\left[\varepsilon_{i t}\right]=E\left[\left(r_{i t}-\alpha_{i}-\beta_{i} r_{m t}\right)\right]=0 \\
& E\left[\varepsilon_{i t} r_{m t}\right]=E\left[\left(r_{i t}-\alpha_{i}-\beta_{i} r_{m t}\right) r_{m t}\right]=0
\end{aligned}
$$

The model is exactly identified. We can "stack" these moment conditions and estimate the parameters $\left\{\alpha_{i}, \beta_{i}\right\}$ of the model, by the usual formulation using sample moments.

The tests discussed in the paper are different ways of testing the parametric restriction $\alpha_{i}=0$.
We will show how the GMM framework can be used to test the CAPM in a couple of examples involving US and Norwegian Data.

## Exercise

Consider the 5 (ew) size based portfolios at the US. Estimate the CAPM using GMM with the Mackinlay method.
Use data 1926-2012.

1. Does the CAPM seem like a sufficient model for pricing this crossection of stock returns?

## Exercise Solution

R code to do analysis

## Read data

source("read_pricing_factors.R")
source("read_size_portfolios.R")
eRi <- FFSize5EW - RF
data <- merge(eRi,RMRF,all=FALSE)
eRi <- as.matrix (data[,1:5])
eRm <- as.matrix(data[,6])

| Index | Lo20 | Qnt2 |  |
| :---: | :---: | :---: | :---: |
| Min. :1926 | Min. : -32.010 | Min. : -31.9600 | Min |
| 1st Qu.:1948 | 1st Qu.: -3.110 | 1st Qu.: -2.8650 | st Q |
| Median :1970 | Median : 0.960 | Median : 1.1700 | Media |
| Mean :1970 | Mean : 1.373 | Mean : 0.9704 | Mean |
| 3rd Qu.:1991 | 3rd Qu.: 4.695 | 3rd Qu.: 4.5350 | 3rd |
| $\begin{gathered} \text { Max. } \quad: 2013 \\ \text { Qnt4 } \end{gathered}$ | Max. : 110.670 Hi20 | Max. : 81.1900 <br> RMRF | Max. |
| Min. : -29.760 | Min. : -30.100 | Min. :-28.980 |  |
| 1st Qu.: -2.470 | 1st Qu.: -2.195 | 1st Qu.: -2.105 |  |
| Median : 1.160 | Median : 0.930 | Median : 1.010 |  |
| Mean : 0.787 | 7 Mean : 0.655 | Mean : 0.628 |  |
| 3rd Qu.: 4.125 | 5 3rd Qu.: 3.640 | 3rd Qu.: 3.655 |  |
| Max. : 50.010 | Max. : 41.790 | Max. : 37.770 |  |

## Exercise Solution

First look at the OLS estimates, low20: (smallest size)

Call:

```
lm(formula = Lo20 ~ eRm)
```

Residuals:

| Min | $1 Q$ | Median | 3Q | Max |
| ---: | ---: | ---: | ---: | ---: |
| -18.245 | -2.920 | -0.618 | 1.890 | 76.656 |

Coefficients:

|  | Estimate | Std. Error | t value $\operatorname{Pr}(>\|t\|)$ |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| (Intercept) | 0.47537 | 0.19852 | 2.395 | $0.0168 *$ |
| eRm | 1.42903 | 0.03633 | 39.334 | $<2 e-16 * * *$ |

Residual standard error: 6.344 on 1033 degrees of freedom Multiple R-squared: 0.5996,Adjusted R-squared: 0.5992
F-statistic: 1547 on 1 and 1033 DF, p-value: < 2.2e-16

## Exercise Solution

|  | 1 | 2 | 3 | 4 | 5 |
| :--- | :---: | :---: | :---: | :---: | :---: |
| (Intercept) | $0.475^{* *}$ | 0.096 | 0.096 | 0.056 | -0.004 |
|  | $(0.199)$ | $(0.077)$ | $(0.077)$ | $(0.051)$ | $(0.029)$ |
| eRm | $1.429^{* * *}$ | $1.260^{* * *}$ | $1.260^{* * *}$ | $1.164^{* * *}$ | $1.049^{* * *}$ |
|  | $(0.036)$ | $(0.014)$ | $(0.014)$ | $(0.009)$ | $(0.005)$ |
| Adj. R | 0.599 | 0.886 | 0.886 | 0.939 | 0.974 |
| Num. obs. | 1035 | 1035 | 1035 | 1035 | 1035 |
| ${ }^{* * *} p<0.01,{ }^{* *} p<0.05,{ }^{*} p<0.1$ |  |  |  |  |  |
|  |  |  |  |  |  |

## Exercise Solution

$$
\operatorname{gmm}(g=e R i \sim e R m, x=e R m)
$$

Coefficients:

|  | Estimate | Std. Error | t value |
| :---: | :---: | :---: | :---: |
| Lo20_ (Intercept) | $4.7537 \mathrm{e}-01$ | $1.8122 \mathrm{e}-01$ | $2.6232 \mathrm{e}+00$ |
| Qnt2_(Intercept) | $1.3007 e-01$ | $1.0484 \mathrm{e}-01$ | $1.2406 \mathrm{e}+00$ |
| Qnt3_(Intercept) | $9.5565 \mathrm{e}-02$ | $7.1679 \mathrm{e}-02$ | $1.3332 \mathrm{e}+00$ |
| Qnt4_ (Intercept) | $5.6185 \mathrm{e}-02$ | $5.1781 \mathrm{e}-02$ | $1.0850 \mathrm{e}+00$ |
| Hi20_(Intercept) | -3.7811e-03 | $2.7477 \mathrm{e}-02$ | -1.3761e-01 |
| Lo20_eRm | $1.4290 \mathrm{e}+00$ | $1.0552 \mathrm{e}-01$ | $1.3543 e+01$ |
| Qnt2_eRm | $1.3382 \mathrm{e}+00$ | $5.9089 \mathrm{e}-02$ | $2.2646 \mathrm{e}+01$ |
| Qnt3_eRm | $1.2601 \mathrm{e}+00$ | $3.6174 \mathrm{e}-02$ | $3.4833 \mathrm{e}+01$ |
| Qnt4_eRm | $1.1637 \mathrm{e}+00$ | $1.9760 \mathrm{e}-02$ | $5.8889 \mathrm{e}+01$ |
| Hi20_eRm | $1.0490 \mathrm{e}+00$ | $1.1175 \mathrm{e}-02$ | $9.3867 e+01$ |
| J-Test: degrees of freedom is 0 |  |  |  |
| Test $\mathrm{E}(\mathrm{g})=0$ : | 1.7729137049673 | -23 ******* |  |

## Exercise Solution

Model 1

| Lo20 (Intercept) | $0.475(0.181)^{* * *}$ |
| :--- | :---: |
| Qnt2 (Intercept) | $0.130(0.105)$ |
| Qnt3 (Intercept) | $0.096(0.072)$ |
| Qnt4 (Intercept) | $0.056(0.052)$ |
| Hi20 (Intercept) | $-0.004(0.027)$ |
| Lo20 eRm | $1.429(0.106)^{* * *}$ |
| Qnt2 eRm | $1.338(0.059)^{* * *}$ |
| Qnt3 eRm | $1.260(0.036)^{* * *}$ |
| Qnt4 eRm | $1.164(0.020)^{* * *}$ |
| Hi20 eRm | $1.049(0.011)^{* * *}$ |
| Criterion function | 0.000 |
| Num. obs. | 1035 |
| ${ }^{* * *} p<0.01,{ }^{* *} p<0.05,{ }^{*} p<0.1$ |  |

## Exercise Solution

Now, let us get to the test of whether the intercept is zero: Need to formulate the linear restrictions in matrix form.
$>R<-\operatorname{cbind}(\operatorname{diag}(5)$, matrix $(0,5,5))$
$>c$ <- rep $(0,5)$
> print(R)

|  | $[, 1]$ | $[, 2]$ | $[, 3]$ | $[, 4]$ | $[, 5]$ | $[, 6]$ | $[, 7]$ | $[, 8]$ | $[, 9]$ | $[, 10]$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $[1]$, | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| $[2]$, | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| $[3]$, | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| $[4]$, | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| $[5]$, | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |

> print(c)
[1] 00000

## Exercise Solution

and then perform the test:
> linearHypothesis(res,R,c,test="F")
Linear hypothesis test

Hypothesis:
Lo20_((Intercept) $=0$
Qnt2_( (Intercept) $=0$
Qnt3_((Intercept) $=0$
Qnt4_( (Intercept) $=0$
Hi20_((Intercept) $=0$

Model 1: restricted model
Model 2: eRi ~ eRm

```
    Df Chisq Pr(>Chisq)
1
2 13.588 0.01845 *
```

Fisher Black, Michael Jensen, and Myron Scholes. The capital asset pricing model, some empirical tests. In Michael C Jensen, editor, Studies in the theory of capital markets. Preager, 1972.
Michael R Gibbons. Multivariate tests of financial models, a new approach. Journal of Financial Economics, 10:3-27, March 1982.
Michael R Gibbons, Stephen A Ross, and Jay Shanken. A test of the efficiency of a given portfolio. Econometrica, 57:1121-1152, 1989.

A Craig MacKinlay and Matthew P Richardson. Using generalized method of moments to test mean-variance efficiency. Journal of Finance, 46:511-27, 1991.

