

Estimating CAPM using GMM



Testing the CAPM in a GMM setting

Consider the CAPM relationship, specified by the "moment condition" $% \left({{{\rm{CAPM}}} \right) = {{\rm{CAPM}}} \right)$

 $E[er_{it}] = \beta_i E[er_{mt}]$

where e_{it} is the excess return on an asset or a portfolio, and e_{mt} the excess return on the market portfolio.

One can alternatively specify the CAPM more generally as

$$E[r_{it}] = E[r_{zt}] + \beta_i (E[r_{mt}] - E[r_{zt}])$$

where we let r_{zt} be the return on a "zero covariance" portfolio.

The typical way of testing this relationship, the Black et al. [1972] method, estimates the corresponding regression

 $er_{it} = \alpha_i + \beta_i er_{mt} + \varepsilon_{it}$

and tests whether $\alpha_i = 0$ on an equation by equation basis. Econometically wasteful: the restriction on the constant will hold for all assets, one want to test this jointly, in a multivariate setting. Doing this was proposed by Gibbons [1982], who showed how one could construct a multivariate statistic for testing this. His method was later expanded upon by Gibbons et al. [1989]. These statistic were developed under distributional assumptions that allowed us to use Maximum Likelihood, namely multivariate normality. What if these are not fulfilled, can we still construct a similar test statistic?

This is done in MacKinlay and Richardson [1991] (MR). They construct a test statistic that essentially tests the same restriction, that $\alpha_I = 0$ or that $\alpha_i = E[r_{zt}](1 - \beta_i)$, but in a GMM framework, not a ML.

The setup: The usual regression

$$r_{it} = \alpha_i + \beta_i r_{mt} + \varepsilon_{it}$$

Assume that

$$E[\varepsilon_{it}|r_{mt}]=0$$

Two moment restrictions for each asset *i*:

$$E[\varepsilon_{it}] = E[(r_{it} - \alpha_i - \beta_i r_{mt})] = 0$$

$$E[\varepsilon_{it}r_{mt}] = E[(r_{it} - \alpha_i - \beta_i r_{mt})r_{mt}] = 0$$

The model is exactly identified. We can "stack" these moment conditions and estimate the parameters $\{\alpha_i, \beta_i\}$ of the model, by the usual formulation using sample moments.

The tests discussed in the paper are different ways of testing the parametric restriction $\alpha_i = 0$.

We will show how the GMM framework can be used to test the CAPM in a couple of examples involving US and Norwegian Data.

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Consider the 5 (ew) size based portfolios at the US. Estimate the CAPM using GMM with the Mackinlay method. Use data 1926-2012.

1. Does the CAPM seem like a sufficient model for pricing this crossection of stock returns?

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```
R code to do analysis
Read data
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```
source("read_pricing_factors.R")
source("read_size_portfolios.R")
eRi <- FFSize5EW - RF
data <- merge(eRi,RMRF,all=FALSE)
eRi <- as.matrix(data[,1:5])
eRm <- as.matrix(data[,6])</pre>
```

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| Index | Lo20 | Qnt2 | C |
|----------------|------------------|-------------------|--------|
| Min. :1926 | Min. :-32.010 | Min. :-31.9600 | Min. |
| 1st Qu.:1948 | 1st Qu.: -3.110 | 1st Qu.: -2.8650 | 1st Qı |
| Median :1970 | Median : 0.960 | Median : 1.1700 | Mediar |
| Mean :1970 | Mean : 1.373 | Mean : 0.9704 | Mean |
| 3rd Qu.:1991 | 3rd Qu.: 4.695 | 3rd Qu.: 4.5350 | 3rd Qı |
| Max. :2013 | Max. :110.670 | Max. : 81.1900 | Max. |
| Qnt4 | Hi20 | RMRF | |
| Min. :-29.76 | 0 Min. :-30.10 | 0 Min. :-28.980 | |
| 1st Qu.: -2.47 | 0 1st Qu.: -2.19 | 5 1st Qu.: -2.105 | |
| Median : 1.16 | 0 Median : 0.93 | 0 Median : 1.010 | |
| Mean : 0.78 | 7 Mean : 0.65 | 5 Mean : 0.628 | |
| 3rd Qu.: 4.12 | 5 3rd Qu.: 3.64 | 0 3rd Qu.: 3.655 | |
| Max. : 50.01 | 0 Max. : 41.79 | 0 Max. : 37.770 | |

```
Exercise Solution
```

```
First look at the OLS estimates,
low20: (smallest size)
Call:
lm(formula = Lo20 ~ eRm)
Residuals:
   Min 1Q Median 3Q
                                 Max
-18.245 -2.920 -0.618 1.890 76.656
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.47537 0.19852 2.395 0.0168 *
       1.42903 0.03633 39.334 <2e-16 ***
eRm
```

Residual standard error: 6.344 on 1033 degrees of freedom Multiple R-squared: 0.5996,Adjusted R-squared: 0.5992 F-statistic: 1547 on 1 and 1033 DF, p-value: < 2.2e-16

| | 1 | 2 | 3 | 4 | 5 |
|---------------------|----------|----------|----------|----------|----------|
| (Intercept) | 0.475** | 0.096 | 0.096 | 0.056 | -0.004 |
| | (0.199) | (0.077) | (0.077) | (0.051) | (0.029) |
| eRm | 1.429*** | 1.260*** | 1.260*** | 1.164*** | 1.049*** |
| | (0.036) | (0.014) | (0.014) | (0.009) | (0.005) |
| Adj. R ² | 0.599 | 0.886 | 0.886 | 0.939 | 0.974 |
| Num. obs. | 1035 | 1035 | 1035 | 1035 | 1035 |

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 $^{***}p < 0.01, \ ^{**}p < 0.05, \ ^{*}p < 0.1$

gmm(g = eRi ~ eRm, x = eRm)

Coefficients:

| | Estimate | Std. Error | t value | Pr(> |
|------------------|-------------|------------|-------------|------|
| Lo20_(Intercept) | 4.7537e-01 | 1.8122e-01 | 2.6232e+00 | 8. |
| Qnt2_(Intercept) | 1.3007e-01 | 1.0484e-01 | 1.2406e+00 | 2. |
| Qnt3_(Intercept) | 9.5565e-02 | 7.1679e-02 | 1.3332e+00 | 1. |
| Qnt4_(Intercept) | 5.6185e-02 | 5.1781e-02 | 1.0850e+00 | 2. |
| Hi20_(Intercept) | -3.7811e-03 | 2.7477e-02 | -1.3761e-01 | 8. |
| Lo20_eRm | 1.4290e+00 | 1.0552e-01 | 1.3543e+01 | 8. |
| Qnt2_eRm | 1.3382e+00 | 5.9089e-02 | 2.2646e+01 | 1.5 |
| Qnt3_eRm | 1.2601e+00 | 3.6174e-02 | 3.4833e+01 | 7.7 |
| Qnt4_eRm | 1.1637e+00 | 1.9760e-02 | 5.8889e+01 | 0. |
| Hi20_eRm | 1.0490e+00 | 1.1175e-02 | 9.3867e+01 | 0. |
| | | | | |

| J-Test: degrees | of freedom is O | |
|-----------------|---------------------|---------|
| | J-test | P-value |
| Test E(g)=0: | 1.7729137049673e-23 | ****** |

| | Model 1 |
|--------------------|------------------------------|
| Lo20 (Intercept) | 0.475 (0.181)*** |
| Qnt2 (Intercept) | 0.130 (0.105) |
| Qnt3 (Intercept) | 0.096 (0.072) |
| Qnt4 (Intercept) | 0.056 (0.052) |
| Hi20 (Intercept) | -0.004 (0.027) |
| Lo20 eRm | 1.429 (0.106)*** |
| Qnt2 eRm | 1.338 (0.059) ^{***} |
| Qnt3 eRm | 1.260 (0.036)*** |
| Qnt4 eRm | 1.164 (0.020)*** |
| Hi20 eRm | 1.049 (0.011)*** |
| Criterion function | 0.000 |
| Num. obs. | 1035 |
| | |

 $^{***}
ho < 0.01, \ ^{**}
ho < 0.05, \ ^{*}
ho < 0.1$

Now, let us get to the test of whether the intercept is zero: Need to formulate the linear restrictions in matrix form.

```
> R <- cbind(diag(5),matrix(0,5,5))</pre>
> c <- rep(0,5)
> print(R)
    [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] [,9] [,10]
[1.]
   1
          0
              0
                  0
                      0
                          0
                              0
                                   0
                                       0
                                            0
[2,] 0 1 0 0 0
                          0
                              0
                                   0
                                       0
                                            0
[3,] 0 0 1 0 0
                          0
                              0
                                   0
                                       0
                                            0
[4,] 0 0 0 1 0 0
                              0
                                   0
                                       0
                                            0
       0
                    1
[5.]
   0
            0
                  0
                          0
                              0
                                   0
                                       0
                                            0
> print(c)
[1] 0 0 0 0 0
```

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and then perform the test:

> linearHypothesis(res,R,c,test="F") Linear hypothesis test

Hypothesis: Lo20_((Intercept) = 0 Qnt2_((Intercept) = 0 Qnt3_((Intercept) = 0 Qnt4_((Intercept) = 0 Hi20_((Intercept) = 0 Model 1: restricted model Model 2: eRi ~ eRm

```
Df Chisq Pr(>Chisq)
1
2 5 13.588 0.01845 *
```

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